Full Credibility for GLM's and GLMM's

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Joint project with Oscar Alberto Quijano Xacur and Jian Tang

José Garrido – Concordia University Full Credibility for GLM's and GLMM's

Warning

This talk contains explicit mathematical content.

Actuarially inspired guidance is advised (Rated AIG).

Since Mowbray (1914, PCAS) up to de Vylder (1985, IME) actuaries have worried about the reliability of statistical premium estimates obtained from the scarce data observed in some risk classes of a segmented portfolio of policies.

An extensive actuarial literature was produced over those 70 years to study this problem.

As actuaries have been adopting pricing models based on GLM's and GLMM's, the credibility question seems to have disappeared.

Here is a partial list of the few papers on the topic: Nelder and Verrall (1997, AB), Schmitter (2004, AB), Antonio and Beirlant (2007, IME), and G. and Zhou (2009, AB).

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- It is distribution free.

For (conditional) loss distributions, given $\Theta = \theta$, from the exponential (dispersion) family:

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• Bühlmann's formula is exact in this case, i.e. the best linear estimator is also the best overall Bayesian estimator.

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Linear credibility formula for the $\vec{\beta}$'s:

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the left side is not in $R(\mathbf{X})$ for all \mathbf{Z} . So, not possible.

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- Alternatively we could find a prior for $\vec{\mu}$ and find the corresponding distribution for $\vec{\beta}$ with the relation $\vec{\mu} = \mathbf{g}^{-1}(\mathbf{X}\vec{\beta})$.
- The case of 1 parameter is an equivalent problem, but for the exponential family.

Conjugate priors for exponential dispersion families

• An exponential dispersion family can be parametrized as

$$p(y; \mu, \lambda) = a(y, \lambda) \exp \left\{-\frac{\lambda}{2}d(y, \mu)\right\},$$

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• Let us consider the prior for μ defined as

$$\pi_{n_0,x_0}(\mu) = A(n_0,x_0)h(\mu)\exp\Big\{-\frac{n_0}{2}d(x_0,\mu)\Big\},\,$$

for some function h.

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- Example: For the gamma distribution, $h(\mu) = \frac{\log(\mu)}{\mu^N}$ gives

$$\mathbb{E}_{\pi_{n_0,x_0}}[\mu] \propto x_0 \left(1 + \frac{1}{a \ln(x_0) + b}\right),$$

which implies a nonlinear credibility formula for μ .

Bibliography

- Antonio, K. and Beirlant, J. (2007) "Actuarial statistics with generalized linear mixed models. *Insurance: Mathematics and Economics*, 40(1), pp.58–76.
- De Vylder, F.E. (1985) "Non-linear regression in credibility theory", *Insurance: Mathematics and Economics*, **4**(3), pp.163–172.
- Garrido, J. and Zhou, J. (2009) "Full credibility with generalized linear and mixed models", *ASTIN Bulletin*, **39**(61).
- Nelder, J.A. and Verrall, R.J. (1997) "Credibility theory and generalized linear models", ASTIN Bulletin, **27**(1).
- Schmitter, H. (2004) "The sample size needed for the calculation of a GKM tariff", ASTIN Bulletin, 34(1), 2004.
- Mowbray, A.H. (1914) "How extensive a payroll exposure is necessary to give a dependable pure premium", *Proc. of the Casualty Actuarial Society*, **1**, pp.24–30.

Thank you for your attention!