Forward Mortality Modelling of Multiple Populations

Craig Blackburn and Michael Sherris

School of Risk and Actuarial Studies, Australian School of Business University of New South Wales, Sydney, Australia ARC Centre of Excellence in Population Ageing Research

The 49th Actuarial Research Conference July 13-16 2014





Motivation

- Similar populations have similar mortality experience.
- Framework for jointly modelling and forecasting similar populations.
- Forward framework model for systematic longevity risk.
- Risk management application:
 - Portfolio of annuitants aged 65.
 - Benefits of a cross population survivor swap (static hedge).
 - Reduction in unexpected portfolio cash flows.
 - Reduction is solvency capital requirements.



Introduction

- Continuous time stochastic framework two approaches:
 - Short rate model Affine Term Structure Model (ATSM)
 - Forward rate model Heath-Jarrow-Morton (HJM)
- Finite dimensional realisation of HJM → State-space model
- ATSMs (state-space) pass qualitative modelling criteria:
 - Biological reasonable forecasts and distribution
 - Robustness to historical sample period
 - Parsimonious
 - Transparency
 - Ease of implementation (closed-form likelihood function)
- Forward mortality model:
 - Model for systematic longevity risk.
 - Ideal for risk management and pricing applications.



Finite Dimensional Realisation

Finite Dimensional Realisation (FDR)

- Link between HJM and state-space model (ATSM).
- Requires a deterministic volatility function.
- Greater flexibility than ATSM (ATSM is a special case).
- Common factor model: standard ATSM.
- Dependent factor model: no closed-form solution in ATSM.
- FDR allows greater flexibility in designing a model.



Finite Dimensional Realisation

Finite Dimensional Realisation (FDR)

For a given deterministic volatility function $\sigma(\tau)$, the forward mortality process

$$d\mu(t,\tau) = \left\{\frac{\partial\mu(t,t+\tau)}{d\tau} + \sigma(\tau)\int_{0}^{\tau}\sigma(s)'ds\right\}dt + \sigma(\tau)dW^{Q}(t)$$
(1)

has a finite dimensional realisation if the volatility function is of the form

$$\sigma(\tau) = C_0 e^{A\tau} B, \qquad (2)$$

then a state-space representation our system is

$$dZ(t) = A^{P}(\theta^{P} - Z(t))dt + BdW^{P}(t), \qquad Z(0) = \psi$$

$$y(t,\tau) = \frac{\int_{0}^{\tau} C(s)ds}{\tau}Z(t) + \frac{\int_{0}^{\tau} \Theta(s)ds}{\tau} + \epsilon(t).$$



Proposed Model Definition

Dependent Factor Model

3-factor model. Interaction between population dependent factors.

$$\sigma_{1}(\tau) = [\sigma_{1}e^{-\delta_{1}\tau}, \sigma_{12}e^{-\delta_{2}\tau}, \sigma_{3}e^{-\delta_{3}\tau}]$$

$$\sigma_{2}(\tau) = [\sigma_{21}e^{-\delta_{1}\tau}, \sigma_{2}e^{-\delta_{2}\tau}, \sigma_{3}e^{-\delta_{3}\tau}]$$

For the volatility function

$$\sigma(\tau) = C_0 e^{A\tau} B, \qquad (4)$$

The finite dimensional realisation is

$$A = -\begin{pmatrix} \delta_{11} & 0 & 0 & 0 & 0 \\ 0 & \delta_{22} & 0 & 0 & 0 \\ 0 & 0 & \delta_{22} & 0 & 0 \\ 0 & 0 & 0 & \delta_{11} & 0 \\ 0 & 0 & 0 & 0 & \delta_{33} \end{pmatrix} \qquad B = \begin{pmatrix} \sigma_{11} & 0 & 0 \\ 0 & \sigma_{12} & 0 \\ 0 & \sigma_{22} & 0 \\ \sigma_{21} & 0 & 0 \\ 0 & 0 & \sigma_{33} \end{pmatrix}.$$

Proposed Model Definition

Dependent Factor Model

The real-world process (5 states). 'Essentially Affine'

$$dZ(t) = A^{P}(\theta^{P} - Z(t))dt + BdW^{P}(t), \qquad Z(0) = \psi$$
(6)

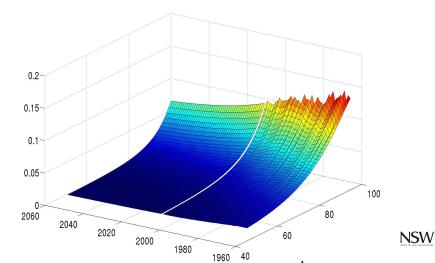
is defined as

$$A^{P} = - \begin{pmatrix} \kappa_{11} & 0 & 0 & 0 & 0 \\ 0 & \kappa_{22} & 0 & 0 & 0 \\ 0 & 0 & \kappa_{33} & 0 & 0 \\ 0 & 0 & 0 & \kappa_{44} & 0 \\ 0 & 0 & 0 & 0 & \kappa_{55} \end{pmatrix},$$

and we set $\theta^{P} = (0, 0, 0, 0, 0)'$.



$y(t, \tau)$ Mortality State-Space Model (Australian Males)

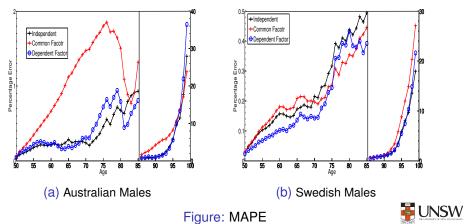


- Fit the state-space model with a standard Kalman filter.
- Historical data from the Human Mortality Database.
 - Australian males ages 50 to 99 for the years 1965 to 2009.
 - Swedish males ages 50 to 99 for the years 1965 to 2009.
- Compare with 2-factor independent model.

Model	Params	Log Likelihood	Aus RMSE	Swe RMSE
Independent Aus.	13	14242.41	0.00126106	-
Independent Swe.	13	15050.73	-	0.00101177
Common Factor	19	27119.67	0.00164847	0.00149174
Dependent Factor	19	29124.43	0.00151793	0.00115905

Table: Fitting Results





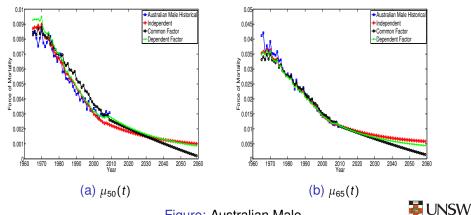


Figure: Australian Male

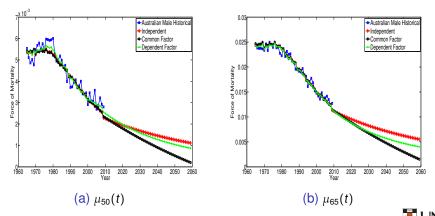


Figure: Swedish Male

- Cohort Survivor Curves

Cohort Survivor Curves

- Deterministic volatility function Gaussian dynamics.
- Volatility is age dependent (not time dependent).
- Martingale measure is not unique.
- Change of measure required:
 - State-space model period mortality rates
 - Forward mortality model cohort mortality rates
- Define *x*₀ the youngest age in the state-space model
- Extract best estimate cohort forecasts from the state-space model.



- Cohort Survivor Curves

Cohort Survivor Curves

a Radon-Nikodym density $\frac{d\widehat{Q}}{d\widehat{Q}}\Big|_{F_t} = e^{-\int_0^t \lambda(s)dW_\mu(s) - \frac{1}{2}\int_0^t |\lambda(s)|^2 ds}$ **b** The mortality process under the cohort measure is $\mu(t, s; x_0) = \left[\mu(0, s; x_0) + \int_0^t \sigma_\mu(u, s; x_0)\lambda(u)du\right] + \int_0^t v_\mu(u, s; x_0)du$ $+ \int_0^t \sigma_\mu(u, s; x_0)dW_\mu^{\widehat{Q}}(u)$ $= \widehat{\mu}(0, s; x_0) + \int_0^t v_\mu(u, s; x_0)du + \int_0^t \sigma_\mu(u, s; x_0)dW_\mu^{\widehat{Q}}(u),$

where $\widehat{\mu}(0, s; x_0)$ is the best estimate cohort mortality rate

$$\widehat{\mu}(\mathbf{0}, \boldsymbol{s}; \boldsymbol{x}_0) = \mu(\mathbf{0}, \boldsymbol{s}; \boldsymbol{x}_0) + \int_0^t \sigma_{\mu}(\boldsymbol{u}, \boldsymbol{s}; \boldsymbol{x}_0) \lambda(\boldsymbol{u}) d\boldsymbol{u}.$$
(7)

Note: no change in the volatility function.

Cohort Survivor Curves

Cohort Survivor Curves

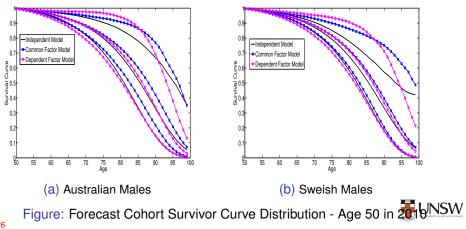
- For a cohort that is aged x at time-t, we know that t - (x - x₀) years ago this cohort was aged x₀.
- We perform the cohort measure change at time- $(t (x x_0))$.
- Requires a change to the volatility function,

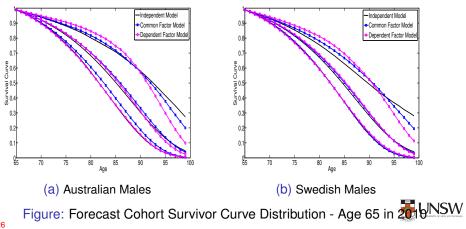
$$\sigma(\tau; \mathbf{x}) = C_0 e^{\mathbf{A}(\tau + (\mathbf{x} - \mathbf{x}_0))} \mathbf{B} = \widehat{C}_0 e^{\mathbf{A}\tau} \mathbf{B},$$
(8)

where $\widehat{C}_0 = C_0 e^{A(x-x_0)}$.

Doesn't effect the estimation of the volatility function.







Longevity Risk Management

- Static hedge of longevity risk with a survivor swap.
- Annuity portfolio of Australian males, aged 65 in 2010.
- Survivor swap based on Australian or Swedish male population (index based).
- Swap fixed for floating annuity payment cash flow.
- No longevity risk premium (Martingale is cohort best estimate)
- Australian survivor swap
 - Systematic longevity risk transfer.
 - Idiosyncratic risk.
- Swedish survivor swap
 - Population basis risk.
 - Idiosyncratic risk.
 - Discrete HJM Monte Carlo simulation.



HJM Simulation

The insurer expected cash flow is

$$S_1(0, t; x_{65}) = e^{-\int_0^t \mu_1(0,s;x_{65})ds}$$

For each simulation path m, the random death time is

$$\tau_i^{(m)} = \inf \left\{ t_i : \sum_{t_s = t_0}^{t_i} \mu_1^{(m)}(t_s; x_{65}) \ge \varrho \right\},\,$$

■ The number of survivors at time-*t_i* for path *m*,

$$\widetilde{S}_{1}^{(m)}(t_{i}; x_{65}) = n_{0} - \widetilde{N}_{1}^{(m)}(t_{i}; x_{65})$$

where,

$$\widetilde{N}_{1}^{(m)}(t_{i}; x_{65}) = \sum_{i=1}^{n_{0}} \mathbf{1}_{\{\tau_{i}^{(m)} \leq t_{i}\}}$$



HJM Simulation

 Unhedged case; the insurer's unexpected cash flow at time-t for path m is

$${}^{u}CF_{1}^{(m)}(t) = S_{1}(0,t;x_{65}) - \widetilde{S}_{1}^{(m)}(t).$$

Survivor swap based on the Australian male population.

$${}^{h}CF_{1}^{(m)}(t) = \widehat{S}_{1}^{(m)}(t) - \widetilde{S}_{1}^{(m)}(t),$$

Survivor swap based on the Swedish male population.

$${}^{h}CF_{2}^{(m)}(t) = S_{1}(0,t;x_{65}) - \widetilde{S}_{1}^{(m)}(t) + \varphi(\widehat{S}_{2}^{(m)}(t) - S_{2}(0,t;x_{65})),$$



Hedge Efficiency - Australian Annuity Portfolio

- The hedge efficiency is given by: $1 \frac{\sigma_{hedged}}{\sigma_{unhedged}}$
- Measure the reduction in unexpected cash flow.
- φ determined from simulation.

	Swap		Portfolio Size		Swap Ratio
	Population	200	1000	10,000	φ
Dependent	Australian Index	56.58%	78.59%	92.94%	1
Factor	Swedish Index	39.67%	53.91%	56.67%	1.1

Table: Hedge Efficiency



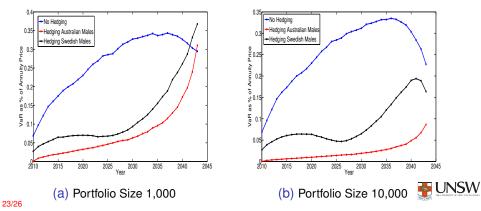
Solvency Capital Requirement

- Insurer remain solvent over a 1-year horizon with 99.5% probability.
- HJM model includes systematic longevity risk.
- Value future liabilities at every time-t.
- Test proposed model with different portfolio sizes.



Solvency Capital Requirement

Dependent factor model



Conclusion

- Present a forward mortality framework for multiple populations.
- Models are fit to historical mortality data.
- Finite dimensional realisation provides a very flexible model.
- Closed-form cohort survivor curves.
- Analyse the hedge efficiency and VaR of a Survivor swap on an annuity portfolio.



Thank you

Questions!



25/26

References

26/26

- [1] Cairns, A. J. G., Blake, D., Dowd, K., Coughlan, G. D., and Khalaf-allah, M. (2011). 'A gravity model of mortality rates for two related populations.' *North American Actuarial Journal*, 15(2):334–356.
- [2] Li, J. S.-H. and Hardy, M. R. (2011). 'Measuring Basis Risk in Longevity Hedges.' North American Actuarial Journal, 15(2):177–200.
- [3] Zhou, R., Wang, Y., Kaufhold, K., Li, J. S.-H., and Tan, K. S. (2012). 'Modeling Mortality of Multiple Populations with Vector Error Correction Models: Applications to Solvency II.' *submitted for publication (2012)*.
- [4] Zhu, N. and Bauer, D. (2011). 'Applications of Forward Mortality Factor Models in Life Insurance Practice*.' The Geneva Papers on Risk and Insurance-Issues and Practice UNSW 36(4):567–594.